Survival of Children With Hypoplastic Left Heart Syndrome

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OBJECTIVE: To examine the survival of infants with hypoplastic left heart syndrome (HLHS) and potential influence of demographic and clinical characteristics on survival using population-based data.

METHODS: Infants with nonsyndromic HLHS (n = 212) born between 1979 and 2005 were identified through the Metropolitan Atlanta Congenital Defects Program. Vital status was ascertained through 2009 based on linkage with vital records. We estimated Kaplan-Meier survival probabilities stratified by select demographic and clinical characteristics.

RESULTS: The overall survival probability to 2009 was 24% and significantly improved over time: from 0% in 1979–1984 to 42% in 1999–2005. Survival probability was 66% during the first week, 27% during the first year of life, and 24% during the first 10 years. Survival of very low and low birth weight or preterm infants and those born in high-poverty neighborhoods was significantly poorer. For children with information on surgical intervention (n = 88), the overall survival was 52%, and preterm infants had significantly poorer survival (31%) compared with term infants (56%). For children who survived to 1 year of age, long-term survival was ~90%.

CONCLUSIONS: Survival to adolescence of children with nonsyndromic HLHS born in metropolitan Atlanta has significantly improved in recent years, with those born full term, with normal birth weight, or in a low-poverty neighborhood having a higher survival probability. Survival beyond infancy to adolescence is high. A better understanding of the growing population of survivors with HLHS is needed to inform resource planning.

WHAT’S KNOWN ON THIS SUBJECT: Hypoplastic left heart syndrome (HLHS) is a critical congenital heart defect with high mortality. With advances in surgical intervention in recent years, survival of infants with HLHS has improved, but information on long-term survival using population-based data is limited.

WHAT THIS STUDY ADDS: In this population-based study, survival to adolescence of children with HLHS has significantly improved in recent years. Among infant survivors, >90% survived up to 18 years. Gestational age, birth weight, and neighborhood poverty may affect survival.
Hypoplastic left heart syndrome (HLHS) is a complex congenital heart defect (CHD) occurring in ~2 per 10,000 livebirths. It is fatal without surgical intervention and responsible for 25% to 40% of all neonatal cardiac mortality. Studies have shown 1-year survival for HLHS ranges from 20% to 60%, with relatively stable 5-year, 10-year, and 15-year survival of ~40%. The highest mortality for infants with HLHS undergoing surgical intervention is usually with stage 1 palliation and the interstage period. Intervention is usually with stage 1 HLHS undergoing surgical intervention is usually with stage 1 palliation and the interstage period between stage 1 and stage 2 surgery, both done in the first year of life. Survival estimates for HLHS have improved in the recent birth era, reflecting significant improvements in management of HLHS. In a national, population-based study using death certificate data from 1979 to 1997, there was no significant decrease in infant mortality attributed to HLHS after the mid-1980s. Another population-based study in Texas showed significantly lower mortality for HLHS in the period from 2001-2003 compared with the period from 1996-2000. A population-based study in New York State linking to in-state mortality records examined the survival of children born between 1983 and 2006 with selected congenital heart defects and found that children with HLHS born in later years had a significantly higher survival probability. However, no other population-based study has examined survival of HLHS beyond childhood using multiple sources of mortality data, across several birth eras, and prognostic factors. In addition to improvements in medical care, other potential prognostic factors include infant characteristics such as low birth weight, prematurity, presence of other birth defects or chromosomal anomalies, race/ethnicity, and high poverty neighborhood. However, the extent to which these factors may influence survival is uncertain.

In this study, we examined changes over time in survival of children with HLHS using population-based birth defects surveillance data collected on children born over a period of 30 years in metropolitan Atlanta, Georgia. In addition, we evaluated clinical and demographic characteristics as potential prognostic factors for such a birth cohort of HLHS.

**METHODS**

**Birth Cohort**

We included children with HLHS ascertained by the Metropolitan Atlanta Congenital Defects Program (MACDP), an ongoing population-based birth defects surveillance system, for birth years 1979 to 2009. MACDP has actively monitored birth defects among live-born infants and fetal deaths of ≥20 weeks’ gestation or >500 g born to resident mothers of the 5-county metropolitan Atlanta area since 1968. MACDP operates in partnership with the Georgia Department of Public Health and has approval of the institutional review board of the Centers for Disease Control and Prevention. Trained abstractors collect demographic and clinical information of structural birth defects, chromosomal abnormalities, and genetic syndromes from several data sources. Additional information about MACDP has been published previously.

The surveillance data for children with HLHS in MACDP were previously reviewed and classified by experts in pediatric cardiology using standard clinical nomenclature from the Society of Thoracic Surgeons and developmental schema. In our study, we included children with isolated HLHS (ie, no other major CHD) and those with associated major extracardiac defects. We defined associated defects as “major” if they had surgical, medical, or serious cosmetic importance. Children with chromosomal abnormalities or known syndromes (eg, trisomies or Noonan syndrome) were excluded.

**Follow-up**

We linked children with HLHS born during the period 1979 to 2005 to Georgia vital records to identify in-state deaths through 2009. We also linked these children with death records available through 2009 in the National Death Index (NDI), a centralized index of death record information compiled by the National Center for Health Statistics. NDI data files were unavailable beyond 2009; hence, we could not trace survival through multiple sources for more recent years. General information and details about the matching process have been described previously. We considered children without death records alive and censored their survival times at the end of the follow-up period on December 31, 2009.

**Prognostic Factors**

In our survival analysis, we considered several demographic and clinical characteristics as potential prognostic factors. Demographic variables included birth period (1979–1984, 1985–1991, 1992–1998, 1999–2005), race/ethnicity (non-Hispanic white, non-Hispanic black, other), infant gender (male, female), maternal age (<20, 20–34; ≥35 years of age), and neighborhood poverty (low, high). Poverty level was based on the percentage of people living in poverty during the closest US census data year (1980, 1990, or 2000) in the census tract linked to the geocoded maternal residence at the time of birth: (1) if <20% of people living in a given census tract were living below the federally defined poverty line, the area was considered as a low-poverty neighborhood; and (2) ≥20%, a high-poverty neighborhood. Using census tract data as proxies for socioeconomic status has been suggested and validated by previous studies. Clinical variables included birth weight (<1500, 1500–2499, ≥2500 g), gestational age (preterm or <37, term or ≥37 weeks), plurality (single, multiple), presence...
of other major extracardiac defects (absent, present), and occurrence of cardiac surgery shortly after birth documented in abstraction record (yes, no). Information on surgical interventions in abstraction records were scarce before 1992; hence, we only included this information on the birth cohorts from 1992 through 2005 and stratified analysis accordingly.

**Statistical Analyses**

We estimated the survival probability of children with HLHS by age and birth cohort using the Kaplan-Meier product-limit method. To calculate the variance of the estimated survival probability and their 95% confidence intervals (CIs), we applied Greenwood’s method. We used the log-rank test to determine whether the survival probabilities were significantly different among different levels of the potential risk factors, both among all children with HLHS and among those who underwent surgical interventions. Variables associated with mortality were included in the Cox proportional hazards regression models. Cox proportional hazard models were used to estimate unadjusted and adjusted hazard ratios (aHR) for possible prognostic factors. The assumption of proportionality for the hazards was checked by plotting estimated log-cumulative hazard versus log of survival time for different categories of the risk factors.

**RESULTS**

Of 229 children with HLHS identified by MACDP born 1979–2005, 212 were classified as nonsyndromic (183 or 86% with no associated major defects) and 17 as having a known chromosomal defect or syndrome. Of the 212 nonsyndromic HLHS children who were the focus of this analysis, 160 died during the follow-up period (23 deaths found only in NDI). The overall survival probability of children with nonsyndromic HLHS was 24%, with most deaths occurring within the first year. The survival probability by age was 95% during the first day, 66% during the first week, 40% during the first month, and 27% during the first year of life, remaining relatively stable thereafter (Table 1). Among children who survived to 1 year of age (n = 58), 90% (n = 52) survived up to 18 years of age. The survival probability varied by birth period and age group (Table 1), being 0% in age groups beyond 1 year in the earliest birth cohort (1979–1984) and 48% among 1- to 5-year-olds for the more recent birth cohort (1999–2005).

<table>
<thead>
<tr>
<th>TABLE 1</th>
<th>Survival Probability by Birth Period and Age Interval Among Children With HLHS, Atlanta, Georgia</th>
</tr>
</thead>
<tbody>
<tr>
<td>Age Interval</td>
<td>All Births 1979–2005</td>
</tr>
<tr>
<td>Birth (&lt;1 y)</td>
<td>No. at Risk</td>
</tr>
<tr>
<td>Birth (&lt;1 y</td>
<td>212</td>
</tr>
<tr>
<td>1 d–&lt;7 d</td>
<td>201</td>
</tr>
<tr>
<td>7 d–&lt;28 d</td>
<td>139</td>
</tr>
<tr>
<td>28 d–&lt;1 y</td>
<td>85</td>
</tr>
<tr>
<td>1 y–&lt;5 y</td>
<td>58</td>
</tr>
<tr>
<td>5 y–&lt;10 y</td>
<td>55</td>
</tr>
<tr>
<td>10 y–&lt;30 y</td>
<td>53</td>
</tr>
<tr>
<td>Overall (&lt;=30 y)</td>
<td>212</td>
</tr>
</tbody>
</table>

Birth period: 1979–2005; follow-up: December 31, 2009; unadjusted. —, no data available beyond 10 years of age for this birth period.
TABLE 2  Survival Probability of Children With HLHS by Selected Demographic and Clinical Characteristics, Atlanta, Georgia

<table>
<thead>
<tr>
<th>Characteristic</th>
<th>No. of Births</th>
<th>No. of Deaths (%)</th>
<th>Survival % (95% CI)</th>
<th>P (Log-Rank Test)</th>
</tr>
</thead>
<tbody>
<tr>
<td>Birth period</td>
<td></td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td>1979–1984</td>
<td>41</td>
<td>41 (100.0)</td>
<td>0.00</td>
<td>&lt;.001</td>
</tr>
<tr>
<td>1985–1991</td>
<td>58</td>
<td>54 (93.1)</td>
<td>6.9 (2.2–15.3)</td>
<td></td>
</tr>
<tr>
<td>1992–1998</td>
<td>53</td>
<td>32 (60.4)</td>
<td>39.3 (26.2–52.2)</td>
<td></td>
</tr>
<tr>
<td>1999–2005</td>
<td>60</td>
<td>33 (55.0)</td>
<td>42.5 (28.9–55.4)</td>
<td></td>
</tr>
<tr>
<td>Race/ethnicitya</td>
<td></td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td>White (non-Hispanic)</td>
<td>104</td>
<td>78 (75.0)</td>
<td>23.9 (15.8–32.8)</td>
<td>.45</td>
</tr>
<tr>
<td>Black (non-Hispanic)</td>
<td>83</td>
<td>66 (79.5)</td>
<td>20.3 (12.4–29.6)</td>
<td></td>
</tr>
<tr>
<td>Other</td>
<td>23</td>
<td>16 (68.8)</td>
<td>30.4 (13.5–48.3)</td>
<td></td>
</tr>
<tr>
<td>Infant genderb</td>
<td></td>
<td></td>
<td></td>
<td>.51</td>
</tr>
<tr>
<td>Male</td>
<td>118</td>
<td>92 (78.0)</td>
<td>21.1 (14.0–29.3)</td>
<td></td>
</tr>
<tr>
<td>Female</td>
<td>94</td>
<td>68 (72.3)</td>
<td>27.0 (18.3–36.4)</td>
<td></td>
</tr>
<tr>
<td>Maternal age (y)c</td>
<td></td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td>&lt;20</td>
<td>28</td>
<td>23 (82.1)</td>
<td>17.9 (6.5–33.7)</td>
<td>.17</td>
</tr>
<tr>
<td>20–34</td>
<td>164</td>
<td>125 (76.2)</td>
<td>22.7 (18.3–29.7)</td>
<td></td>
</tr>
<tr>
<td>≥35</td>
<td>20</td>
<td>12 (60.0)</td>
<td>40.0 (19.3–60.0)</td>
<td></td>
</tr>
<tr>
<td>Birth weight (g)</td>
<td></td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td>&lt;1500</td>
<td>4</td>
<td>4 (100.0)</td>
<td>0.0</td>
<td>.01</td>
</tr>
<tr>
<td>1500–2499</td>
<td>32</td>
<td>27 (84.4)</td>
<td>15.6 (5.7–30.0)</td>
<td></td>
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<tr>
<td>≥2500</td>
<td>176</td>
<td>129 (73.3)</td>
<td>25.6 (19.1–32.6)</td>
<td></td>
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<tr>
<td>Gestational age (wk)b</td>
<td></td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td>Preterm (&lt;37)</td>
<td>34</td>
<td>30 (88.2)</td>
<td>11.8 (3.7–24.9)</td>
<td>.02</td>
</tr>
<tr>
<td>Term (≥7)</td>
<td>177</td>
<td>129 (72.9)</td>
<td>26.0 (19.5–33.0)</td>
<td></td>
</tr>
<tr>
<td>Plurality</td>
<td></td>
<td></td>
<td></td>
<td>.19</td>
</tr>
<tr>
<td>Single</td>
<td>202</td>
<td>151 (74.8)</td>
<td>24.3 (18.4–30.7)</td>
<td></td>
</tr>
<tr>
<td>Multiple</td>
<td>10</td>
<td>9 (90.0)</td>
<td>10.0 (6.6–35.8)</td>
<td></td>
</tr>
<tr>
<td>Other major extracardiac defects</td>
<td></td>
<td></td>
<td></td>
<td>.38</td>
</tr>
<tr>
<td>Absent</td>
<td>183</td>
<td>139 (76.0)</td>
<td>23.1 (17.1–28.8)</td>
<td></td>
</tr>
<tr>
<td>Present</td>
<td>29</td>
<td>21 (72.4)</td>
<td>27.8 (13.1–44.3)</td>
<td></td>
</tr>
<tr>
<td>Neighborhood povertya</td>
<td></td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td>Low</td>
<td>173</td>
<td>127 (73.4)</td>
<td>25.4 (18.9–32.5)</td>
<td>.01</td>
</tr>
<tr>
<td>High</td>
<td>32</td>
<td>29 (90.6)</td>
<td>9.4 (2.4–22.3)</td>
<td></td>
</tr>
<tr>
<td>Cardiac surgeryb</td>
<td></td>
<td></td>
<td></td>
<td>&lt;.001</td>
</tr>
<tr>
<td>No</td>
<td>23</td>
<td>23 (100.0)</td>
<td>0.0</td>
<td></td>
</tr>
<tr>
<td>Yes</td>
<td>88</td>
<td>40 (45.5)</td>
<td>52.1 (40.1–62.8)</td>
<td></td>
</tr>
</tbody>
</table>

Birth period 1979–2005, follow-up December 31, 2009; unadjusted. Low-poverty neighborhood: <20% of mothers living in a given census tract were living in poverty; high-poverty neighborhood: ≥20% of mothers living in a given census tract were living in poverty.  
a Subjects with unknown status were deleted.  
b 1992–2005 births only; n = 111 with known surgery status.  
c No survivors with the characteristic in the birth period.

DISCUSSION

This population-based survival study of children born with nonsyndromic HLHS over a 30-year period in metropolitan Atlanta shows significant improvement in survival over time. Although overall survival was only 24%, among children born during the more recent birth period of 1992–2005 who underwent surgery, the survival probability was 52%. The majority of mortality occurs in the first year of life, with relatively stable survival after 1 year of age. The survival probability of children with HLHS improved by birth period in all age groups. This finding is consistent with a study in metropolitan Atlanta showing improved 1-year survival over birth era (from 67% in 1979–1993% to 83% in 1994–2005) for infants with several types of critical CHDs.26 Factors found to impair survival included preterm birth, having a very low or low birth weight, or being born to mothers residing in high-poverty neighborhoods.

Our findings of a 1-year survival probability of 27% among all birth cohorts, which includes all children with HLHS regardless of their surgical intervention status, and of 58% among the birth cohort of 1992–2005 undergoing surgery fall within the wide range of infant survival probabilities of 20% to 60% published elsewhere.3–9,13,26 Variation in study survival rates, which we could not be assessed with our data, could be due to differences in operative techniques, different institution types where surgery is done, inclusion of birth eras with limited potential for surgical intervention, or inclusion of infants with chromosomal anomalies. However, although all studies note stabilization in survival beyond infancy, our study demonstrates relative stability through adolescence. We did not have data on longer term outcomes, such as potential heart failure in young adulthood. Among children with HLHS who survived infancy, ~90% survived up

Survival by Surgical Intervention

Limiting our analysis to the 1992–2005 birth period for which there was more reliable information on cardiac surgery than in previous years, survival probability among infants undergoing surgical intervention (n = 88) was 95% during the first week, 78% during the first month, and 58% during the first year of life (Table 3). Among children undergoing cardiac surgery who survived to 1 year of age (n = 51), 94% (n = 48) survived up to 18 years of age. The analysis of the same individual demographic and clinical factors as for the larger group (data not shown) revealed that preterm infants had a significantly poorer overall survival (31%) compared with term infants (56%; P = .02). None of the other factors examined showed a significant effect on survival. Although children who underwent surgical intervention and had normal birth weight or were born in low-poverty neighborhood had a better survival (54% and 61%, respectively) than children with low birth weight or born in high-poverty neighborhood (38% and 47%, respectively), the differences were not statistically significant.
to 18 years of age, a finding similar to
the 10-year survival rate of 81% to
86% observed in a Texas study.7 This
finding of a high survival rate beyond
infancy indicates that the most critical
and vulnerable period of mortality is
the first year of life, and both health
care professionals and parents face
challenging treatment decisions.27–29

With more families opting for
surgery,30 our finding of a high survival
rate beyond infancy also suggests that
a growing number of children with
HLHS are likely to reach adolescence
and adulthood. In a recent population-
based study of pooled data on survival
of children with several types of birth
defects,9 children with HLHS had
similar survival (50%) to 8 years of age
as in our current findings, but much
lower survival compared with other
noncardiac defects, likely due to the
severity of the HLHS malformation.
However, CHDs as a group are
relatively common birth defects, and it
will be important to monitor this
growing population of persons with
HLHS for their long-term care needs
and to better inform planning and
allocation of resources.15,31,32

Our findings on prognostic factors are
consistent with those of a population-
based study in Michigan,14 which
reported that infants with HLHS had
significantly poorer survival at 1 year
of age if they were born preterm, had
low birth weight, and lived in an area
with high poverty. Chromosomal
anomalies such as trisomy 18 and
trisomy 13 are also known to affect
overall survival and mortality.14,15,33 In
our study, we excluded children with
HLHS and chromosomal anomalies so
as to have a more homogenous cohort
in which to examine nongenetic
prognostic factors. The poorer survival
of children born in high-poverty
neighborhoods, a proxy for parental
socioeconomic status, is also
noteworthy and may suggest that
parents may have had more limited
access to care when their children were
born. Some studies have suggested
racial/ethnic disparities in survival for
CHDs including HLHS,6,9,12–14 which
needs to be looked at more closely in
future studies. Possible reasons that we
did not find these disparities could be
small sample size, and inability to
distinguish some racial/ethnic groups.
For example, our data did not
distinguish Hispanics until the 1990s,
and the 80% of the population was
classified as white, black, or other.

Fixler and colleagues had slightly
different findings from the current
study,6 possibly due to differences in
study design, population, and birth
period. They studied mortality up to
age 5 in children with single ventricle
defects born 1996–2003, and
reported that children with HLHS had
42% overall survival at 1 year of age
and 38% overall survival at 5 years of
age. The survival of infants with HLHS
and extracardiac defects was lower at
1 and 5 years of age (27% and 21%)
compared with those without
extracardiac defects (44% and 41%,
respectively). In our study, presence
of other major extracardiac defects
did not affect survival probabilities.
Possible reasons for this finding are
a smaller sample size and less severe
extracardiac defects after exclusion of
chromosomal disorders.

Strengths
We included all live births from
a defined population captured by
a well-established, population-based
birth defects surveillance system,
with active and multiple source
ascertainment methods. Identification
of children with HLHS did not rely solely on administrative coding but resulted from a clinically classified data set, which minimized the potential for misclassification of HLHS. We were able to estimate long-term survival probability based on up to 30 years of follow-up for the earliest birth cohorts, which is longer than has been previously reported. For birth cohorts, which is longer than has been previously reported. For ascertainment of vital status, we used multiple sources, which previously have been shown to be reliable in identifying a high proportion of deaths.34,35 Finally, we were able to examine a wide range of demographic and clinical characteristics as potential predictors of survival.

Limitations

First, we assumed that children with HLHS for whom no date of death was available were alive at the end of follow-up. Although it is possible that some deaths might have been missed because of missing data on a linkage variable, it is likely that the number of missed deaths was small, because we used multiple data sources and the NDI uses several factors for linkage.20 A related limitation is that our multisource follow-up was only through 2009. NDI data were not available to search for more recent deaths; further linkage to Georgia vital records through 2013 did not identify any additional in-state deaths. This information on follow-up beyond 2009 is not included in our results because only one data source was used. Second, statistical power to analyze some characteristics as prognostic factors (eg, older maternal age, presence of major extracardiac defects) may have been limited due to the small number of affected individuals with HLHS having such characteristics. Third, because we did not have information on direct indicators of individual socioeconomic status, we had to use a proxy for socioeconomic status, namely, neighborhood poverty indicator based on the mother’s census tract at the time of delivery. We did not have data with which to assess the validity of this proxy as a measure of socioeconomic status. Furthermore, we lacked complete information on other possible demographic and clinical characteristics that could affect survival such as maternal and paternal education and occupation, insurance status, or access to care; therefore, we could not explore the potential impact of these factors on survival. Fourth, our data on surgical intervention was limited to the abstracted information; we did not review original medical records for specific surgical procedures or outcomes. Thus, our reported results may be subject to an underreporting of surgical information.

CONCLUSIONS

In this population-based study, survival to adolescence among children born from 1979 to 2005 with nonsyndromic HLHS in metropolitan Atlanta has improved significantly over time, with those born full-term or undergoing surgery having better survival. Of those who survive infancy, the long-term survival rate was ∼90%. These findings support the concept that with improved treatment, there is a growing population of adults with CHDs, including HLHS. Understanding predictors of survival is one step toward caring for this new aging population and adds to the growing literature on resource use and outcomes for persons with CHDs. Better understanding of the growing population with HLHS and their health care needs can inform future resource planning and allocation. Further work is warranted to assess the potential effect of socioeconomic status and access to care factors on survival of children with HLHS, which may further inform strategies to improve health outcomes.

ACKNOWLEDGMENTS

The authors acknowledge the MACDP staff members for their conscientious and skilled data collection.

ABBREVIATIONS

aHR: adjusted hazard ratio
CHD: congenital heart defect
CI: confidence interval
HLHS: hypoplastic left heart syndrome
MACDP: Metropolitan Atlanta Congenital Defects Program
NDI: National Death Index

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