Development of Heart and Respiratory Rate Percentile Curves for Hospitalized Children

WHAT’S KNOWN ON THIS SUBJECT: Accurately identifying ill hospitalized children with vital signs concerning for clinical deterioration is fundamental to inpatient pediatrics. Normal vital sign ranges for healthy children are useful for outpatient practice but have limited application to detecting deterioration in the hospital setting.

WHAT THIS STUDY ADDS: Percentile curves for heart and respiratory rate in hospitalized children were developed and validated. The distributions differed from existing reference ranges and early warning scores. They may be useful to identify vital signs deviating from ranges expected among hospitalized children.

abstract

OBJECTIVE: To develop and validate heart and respiratory rate percentile curves for hospitalized children and compare their vital sign distributions to textbook reference ranges and pediatric early warning score (EWS) parameters.

METHODS: For this cross-sectional study, we used 6 months of nurse-documented heart and respiratory rates from the electronic records of 14 014 children on general medical and surgical wards at 2 tertiary-care children’s hospitals. We developed percentile curves using generalized additive models for location, scale, and shape with 67% of the patients and validated the curves with the remaining 33%. We then determined the proportion of observations that deviated from textbook reference ranges and EWS parameters.

RESULTS: We used 116 383 heart rate and 116 383 respiratory rate values to develop and validate the percentile curves. Up to 54% of heart rate observations and up to 40% of respiratory rate observations in our sample were outside textbook reference ranges. Up to 38% of heart rate observations and up to 30% of respiratory rate observations in our sample would have resulted in increased EWSs.

CONCLUSIONS: A high proportion of vital signs among hospitalized children would be considered out of range according to existing reference ranges and pediatric EWSs. The percentiles we derived may serve as useful references for clinicians and could be used to inform the development of evidence-based vital sign parameters for physiologic monitor alarms, inpatient electronic health record vital sign alerts, medical emergency team calling criteria, and EWSs. Pediatrics 2013;131:e1150–e1157

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KEY WORDS: percentile curves, early warning score, heart rate, hospital rapid response team, medical emergency team, physiologic monitoring, reference values, respiratory rate

ABBREVIATIONS

BCPE—Box-Cox power exponential
CCHMC—Cincinnati Children’s Hospital Medical Center
CHOP—The Children’s Hospital of Philadelphia
EHR—electronic health record
EWS—early warning score
GAMLSS—generalized additive models for location, scale, and shape
HR—heart rate
MET—medical emergency team
RR—respiratory rate
RRSs—rapid response systems

Dr Bonafide conceptualized and designed the study, acquired the data from The Children’s Hospital of Philadelphia, participated in the analysis and interpretation of the data, and drafted the initial manuscript; Dr Brady conceptualized and designed the study, acquired the data from Cincinnati Children’s Hospital Medical Center, and participated in the analysis and interpretation of the data; Dr Keren conceptualized and designed the study and participated in the analysis and interpretation of the data; Dr Conway participated in the statistical analysis, and participated in the analysis and interpretation of the data; Dr Marsolo acquired the data from Cincinnati Children’s Hospital Medical Center; Dr Conway participated in the statistical analysis, and participated in the analysis and interpretation of the data; and all authors critically revised the manuscript for important intellectual content, and approved the final manuscript as submitted.

(Continued on last page)
The widespread implementation of rapid response systems (RRSs) over the past decade has called attention to the importance of recognizing and responding to early signs of clinical deterioration. RRSs aim to reduce mortality, cardiac arrest, and respiratory arrest outside of ICUs and include afferent and efferent limbs. The afferent limb of RRSs includes tools to aid clinicians in detecting patients exhibiting early signs of clinical deterioration. The efferent limb consists of medical emergency teams (METs) that can be urgently summoned to the bedside to assist in management. Accurately interpreting vital signs is critical to the success of RRSs. To facilitate the recognition of vital signs consistent with clinical deterioration, clinicians may rely on several afferent tools, including monitor alarms, electronic health record (EHR) vital sign alerts, MET calling criteria, and early warning scores (EWSs) that assign point values to vital signs outside of age-based ranges.

Limited data are available to inform the development of evidence-based parameters for monitor alarms, EHR alerts, MET calling criteria, and EWSs. Currently available sources to inform the configuration of these parameters include textbook reference ranges and existing EWSs with parameters developed using consensus opinion. No previous studies have described the actual distributions of heart rate (HR) or respiratory rate (RR), 2 vital signs critical to ongoing surveillance for deterioration, among hospitalized children. In this study, we sought to develop and validate HR and RR percentile curves for pediatric inpatients. We aimed to determine the proportion of vital signs in our cohort that fell outside textbook reference ranges and those that would be assigned points using existing EWSs. Finally, we aimed to directly compare our HR and RR percentile cut points to those from a recent meta-analysis of vital sign distributions in well children.

METHODS

Data Sources

We performed this cross-sectional study among children <18 years of age hospitalized on general medical and surgical wards in 2 tertiary-care children's hospitals: Cincinnati Children's Hospital Medical Center (CCHMC) and The Children's Hospital of Philadelphia (CHOP). In both settings, nurses or nursing assistants entered vital signs into discrete fields on EHR flowsheets; CCHMC used the Integrating Clinical Information System (Siemens Medical Solutions, Malvern, PA) and CHOP used Epic Systems (Verona, WI) during the study periods. The minimum frequency of vital sign assessment was determined by physicians' orders. The procedures used to measure HR and RR were neither standardized across patients nor documented in the record. We extracted vital signs, demographics, and primary discharge diagnosis from CCHMC between July 1, 2008, and December 31, 2008, and CHOP between February 1, 2011, and July 30, 2011. The use of distinct time periods allowed us to create a merged data set including each of the 4 seasons, as well as both geographic and temporal variation among the included patients.

Data Quality

We implemented several data integrity measures to ensure that the data were of the highest quality possible. We excluded admissions in which the primary discharge diagnosis was missing. We included only those observations in which the HR and RR were entered simultaneously. We excluded observations in which the RR exceeded the HR or the HR or RR included a decimal. In addition, we excluded observations in which either vital sign met criteria for physiologic implausibility, defined by consensus of the investigators as being more likely to represent keystroke errors than actual observations (HR >300 or <30, and RR >120 or <5).

Sampling Strategy

To minimize ascertainment bias from vital signs collected frequently over short periods of time, we first divided each admission into 6-hour intervals. We then randomly selected 1 HR and 1 RR from within each interval. To reduce ascertainment bias from single individuals with lengthy hospitalizations, we included a maximum of 10 intervals per admission. We then randomly divided the data into a curve development set made up of 67% of the patients and a curve validation set made up of 33%.

Data Analysis

Before developing the curves, we log-transformed the highly skewed RR data to more closely approximate a normal distribution. We also noted evidence of digit preference. For example, among those aged >15 years, 73% of RRs were accounted for by the values 16, 18, and 20. We therefore added uniformly distributed random noise to the RR data before developing the curves to facilitate identifying the underlying distribution of RR in older children. Adding random noise, or “jitter,” can be used to identify the underlying distribution in data that have been subject to rounding or digit preference.

We then developed percentile curves with the vital sign on the y-axis and the patient age at the time the vital sign was measured on the x-axis using the Box-Cox power exponential (BCPE) distribution in the generalized additive models for location, scale, and shape (GAMLSS) package for R software. The use of GAMLSS models with the BCPE distribution in the development of percentile curves for growth
has been demonstrated25,29–31; GAMLS models can readily be applied to vital sign curves. BCPE is an appropriate distribution because it is useful for data exhibiting both skewness and kurtosis, and it allows for determination of z-scores at any point in the distribution. This feature facilitates future development of EWSs that use z-scores to standardize the assignment of point values to vital signs across age groups.

Next, we determined the stability of the first and 99th percentiles by creating curves using 3 bootstrapped samples for both HR and RR and comparing the differences between the 3 bootstrapped samples.

To validate the curves, we divided the validation set into 5 age groups and determined the proportion of vital signs across age groups. To evaluate the impact of respiratory disease on the RR and HR curves, we performed a sensitivity analysis. We identified all admissions with a primary discharge diagnosis of a Disease of the Respiratory System as defined by the International Classification of Diseases, Ninth Revision, Clinical Modification, codes 460 through 519. We then excluded those admissions from the data set. We compared the curves excluding respiratory disease to the curves without disease exclusions.

To evaluate the clinical relevance of our findings to the care of hospitalized children, we determined the proportion of observations deviating from reference ranges that hospitals may use to identify vital signs consistent with deterioration, set monitor alarm parameters, or configure EHR vital sign alerts. We included 3 commonly used pediatric references: Nelson Textbook of Pediatrics,18 The Harriet Lane Handbook,10,20 and the American Heart Association Pediatric Advanced Life Support Provider Manual.21 We also determined the proportion of observations that would have resulted in increased pediatric EWSs by using 2 published scores that are sufficiently described to permit retrospective scoring: Parshuram’s Bedside Pediatric Early Warning System Score17 and Akre’s Pediatric Early Warning Score.16 Of note, Akre’s score was modified from a score previously described by Monaghan10 and Tucker15; however, neither Monaghan nor Tucker’s published description included vital sign ranges.

Because we used a deidentified data set for the analysis, this project was granted exemptions from the CHOP and CCHMC Institutional Review Boards.

RESULTS

We extracted data from 11,028 admissions among 8894 patients aged <18 years at CHOP and 6271 admissions among 5208 patients aged <18 years at CCHMC. Before implementing the sampling strategy described earlier, scatterplots of the data showed numerous “spikes” in the HR and RR distributions, many of which represented clusters of extreme values in individual patients with repeated measurement over short time periods, extreme values in individual patients with repeated measurement across lengthy admissions, or both. We minimized the presence of these spikes by applying the aforementioned sampling strategy, after which 116,383 HR values and 116,383 RR values from 14,014 patients across 17,153 admissions remained (3139, or 18% were repeat admissions). This sample represented 23% of the observations across 99% of the patients in the original data set. Demographics are shown in Tables 1 and 2.

We then used the 77,825 HR and 77,825 RR values in the development set to generate HR and RR percentile curves, shown in Fig 1. Tables 3 and 4 show our suggested clinical cut points at the first, fifth, 10th, 50th, 90th, 95th, and 99th percentiles for HR and RR, respectively, among the 13 age groups. In evaluating the stability of the first and 99th percentiles, we found that the

### Table 1 Patient Characteristics

<table>
<thead>
<tr>
<th>Characteristic</th>
<th>Patients, n (%)</th>
</tr>
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<tbody>
<tr>
<td><strong>Race/ethnicity</strong></td>
<td></td>
</tr>
<tr>
<td>White</td>
<td>7558 (53.9)</td>
</tr>
<tr>
<td>African American</td>
<td>4206 (30.0)</td>
</tr>
<tr>
<td>Hispanic</td>
<td>604 (4.3)</td>
</tr>
<tr>
<td>Asian/Pacific Islander</td>
<td>305 (2.2)</td>
</tr>
<tr>
<td>American Indian/Alaska Native</td>
<td>15 (0.1)</td>
</tr>
<tr>
<td>Multiracial</td>
<td>1 (&lt;0.1)</td>
</tr>
<tr>
<td>Other</td>
<td>1311 (9.4)</td>
</tr>
<tr>
<td>Unknown</td>
<td>14 (0.1)</td>
</tr>
<tr>
<td><strong>Total</strong></td>
<td>14,014</td>
</tr>
<tr>
<td><strong>Gender</strong></td>
<td></td>
</tr>
<tr>
<td>Male</td>
<td>7539 (53.8)</td>
</tr>
<tr>
<td>Female</td>
<td>6474 (46.2)</td>
</tr>
<tr>
<td>Unknown</td>
<td>1 (&lt;0.1)</td>
</tr>
<tr>
<td><strong>Total</strong></td>
<td>14,014</td>
</tr>
</tbody>
</table>

*Self-reported by patient or family and categorized according to the options available in the electronic health record.

### Table 2 Observations by Age Groups

<table>
<thead>
<tr>
<th>Age Group</th>
<th>Observations, n (%)</th>
</tr>
</thead>
<tbody>
<tr>
<td>0–&lt;3 mo</td>
<td>9872 (8.5)</td>
</tr>
<tr>
<td>3–&lt;6 mo</td>
<td>5604 (4.8)</td>
</tr>
<tr>
<td>6–&lt;9 mo</td>
<td>4248 (3.7)</td>
</tr>
<tr>
<td>9–&lt;12 mo</td>
<td>3761 (3.2)</td>
</tr>
<tr>
<td>12–&lt;18 mo</td>
<td>6582 (5.7)</td>
</tr>
<tr>
<td>18–&lt;24 mo</td>
<td>4957 (4.2)</td>
</tr>
<tr>
<td>2–&lt;3 y</td>
<td>8624 (7.7)</td>
</tr>
<tr>
<td>3–&lt;4 y</td>
<td>6585 (5.7)</td>
</tr>
<tr>
<td>4–&lt;6 y</td>
<td>10,896 (9.4)</td>
</tr>
<tr>
<td>6–&lt;8 y</td>
<td>8115 (7.5)</td>
</tr>
<tr>
<td>8–&lt;12 y</td>
<td>15,428 (13.5)</td>
</tr>
<tr>
<td>12–&lt;15 y</td>
<td>14,477 (12.4)</td>
</tr>
<tr>
<td>15–&lt;18 y</td>
<td>16,244 (14.0)</td>
</tr>
<tr>
<td><strong>Total</strong></td>
<td>116,383</td>
</tr>
</tbody>
</table>

*Because age at time of vital sign measurement changes during the course of the hospitalization, these data are presented in terms of the number of observations.
bootstrapped samples resulted in a maximum change in the first and 99th percentiles of 1 beat per minute for HR and 1 breath per minute for RR.

In validating the curves using the remaining 38,558 observations for each vital sign, we found that the proportion of vital signs below the fifth and above the 95th percentile lines for each age group in the validation set was within 0.9% of the expected 5% for the HR data at all ages. However, the proportion of RR observations in the validation set below the fifth and above the 95th percentile differed from the expected 5% by up to 4.7%; the proportion of RRs below the fifth percentile for RR was 9.7% among 12- to 18-year-olds. Because the majority of RRs in older children are accounted for by a small number of unique values, any distribution would likely result in some clustering at the most common values, precluding a truly normal distribution of RR.

We then evaluated the impact of our introduction of uniformly distributed random noise to the RR data as a means of overcoming digit preference. Using the validation data set, we analyzed the distribution with and without random noise added. We found that adding random noise caused minimal change in the distribution of the data (z-scores had a mean of 0.00 and SD of 0.93 for the data set without random noise, and a mean of 0.01 and SD of 0.98 for the data set with random noise) while improving the fit of the curves at the outer percentiles (the proportion of RRs below the fifth and above the 95th percentile lines for each age group in the validation set was within 1.4% of the expected proportion of 5%).

In the sensitivity analysis (shown as dashed lines in Fig 1), we excluded 3713 (21.6%) admissions with primary diagnoses of “diseases of the respiratory system,” leaving 91,763 HR and 91,763 RR values. We found the impact of excluding respiratory disease to be most pronounced at the 99th percentile lines. For HR, excluding resulted in a reduction in the 99th percentile by up to 8 beats per minute, occurring between 5 and 6 years of age. For RR, excluding resulted in a reduction in the 99th percentile by up to 7 breaths per minute, occurring between 4 and 5 years of age.

For each of the reference ranges and EWSs described, we determined the proportion of observations that either deviated from the reference range or earned points in the EWS. We found that 12% to 54% of HR observations and 32% to 40% of RR observations deviated from textbook ranges (Fig 2). In addition, 14% to 38% of HR observations and 15% to 30% of RR observations would have resulted in increased total EWSs (Fig 3). Compared with the data from a recent meta-analysis of HR and RR among well children, the curves we developed by using hospitalized children had a wider span from first to 99th percentile, primarily because of higher HR and RR at the 99th percentile in all age groups (Supplemental Fig 1).

DISCUSSION

Our percentile curves, developed by using a large multicenter sample, provide the first evidence-based reference ranges for HR and RR among hospitalized children. These ranges differed from textbook reference ranges, EWS parameters created by using expert consensus, and distributions of vital signs from well children.
The ranges we developed represent the distributions of HR and RR among ill children hospitalized outside of ICU settings. By using these data, we found that up to 54% of vital sign values among patients in this study were abnormal according to textbook reference ranges, and up to 38% would be assigned points using EWSs with parameters developed based on the consensus of experts. Our results provide the most useful evidence to date for helping clinicians determine the degree to which vital signs deviate from the expected value for hospitalized children. These findings may serve as useful clinical references and can inform the development of evidence-based vital sign parameters for physiologic monitor alarms, EHR vital sign alerts, MET calling criteria, and EWSs.

A recent meta-analysis reviewed 69 studies with HR data for 143,346 children and RR data for 3881 well children in nonhospital settings and used the summary statistics from each study to derive percentile curves. While the meta-analysis provides the best available data to inform out-patient vital sign reference ranges, the uppermost percentiles for HR and RR in our study were higher than in the meta-analysis, raising questions about the application of vital sign distributions from well patients to inform deterioration surveillance in the hospital.

Developing valid vital sign-based tools to detect early signs of clinical deterioration with high sensitivity and low false-positive rates (high specificity) is challenging. Because children have vital sign ranges that vary by age, tool development in pediatrics is particularly complex. Several pediatric EWSs have been described with parameters developed using the consensus of experts. Parshuram's Bedside Pediatric Early Warning System Score (adapted from Monaghan's) and Akre's Pediatric Early Warning Score are the 2 most rigorously studied. Parshuram's score had a sensitivity of 64% and a specificity of 94% in a validation study. Akre's score had a sensitivity of 86%, its specificity was not reported. Neither has been evaluated in a clinical trial. The reference ranges for hospitalized children in our study provide an opportunity to take a data-driven approach to optimizing EWS parameters with the goal of increasing the sensitivity and specificity of these tools. The z-scores derived from the vital sign data could be used to assign EWS point values based on the degree to which the HR or RR deviates from the expected value for age.

In addition to EWSs, the percentiles derived in this study have potential application to configuring physiologic monitor alarm parameters and EHR vital sign alerts. Age-based monitor alarm thresholds based on clinical data from hospitalized children may help minimize alarm fatigue by safely reducing the frequency of false alarms that do not identify meaningful departures from expected vital signs in the hospital. For example, upon hospital admission, HR alarm thresholds could be set at the fifth and 95th percentiles for age. Similarly, inpatient EHR alerts, which may range from highlighting the text of an out-of-range vital sign entry in an electronic flowsheet to generating an alert if a vital sign exceeds an extreme threshold, could be set based on the percentiles. Our study has several limitations. First, we derived these curves by using data from hospitalized children.
from 2 tertiary-care children’s hospitals. The vital sign distributions from patients in community hospitals may differ, limiting the generalizability of our reference ranges to those settings. Second, developing percentile-based EWSs and alarm thresholds assumes that the patients at highest risk of deterioration are those with the most statistically abnormal vital signs, which is likely an oversimplification.33,34 Clinical trials are needed to determine whether percentile-based EWSs have a role in identifying critical but reversible disease states and improving patient outcomes. Third, although we noted that developing EWSs based on normal vital signs from healthy children may result in tools with low specificity and high false-positive rates, developing EWSs based on vital signs from ill children may result in tools with low sensitivity and high false-negative rates. Fourth, we used vital sign data that was manually entered in the course of clinical care. This introduced the issue of digit preference. This issue could likely be overcome either by using research staff to measure vital signs by counting HR and RR for at least 1 full minute or by using physiologic monitor data. We chose observations documented in the course of clinical care because our curves will be referenced in clinical settings, and we aimed for our distributions to be representative of that context. We chose manually entered data over data from monitors because only a fraction of inpatients at these hospitals are monitored, thus using monitor data would introduce

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**FIGURE 2**

Scatterplot array showing the distribution of HR and RR in the study sample compared with textbook reference ranges. Each point on the scatterplot represents 1 vital sign observation. For each reference range, observations that would be considered normal are colored black, and observations that would be considered abnormal are colored red. We found that 12% to 54% of HR observations and 32% to 40% of RR observations in our study sample deviated from the ranges provided.
selection bias. Fifth, just as within-subject changes in height, weight, and head circumference over time are critical to identifying abnormal growth, we suspect that within-subject changes in vital signs over time may be equally if not more important than exceeding absolute percentiles. To address this issue, the z-scores derived from these curves could be used to standardize the degree of within-subject change across age groups in an EWS.

CONCLUSIONS

In summary, we developed and validated the first evidence-based percentile curves for HR and RR among children hospitalized on general wards at tertiary-care pediatric hospitals. Their vital sign distributions differed from existing reference ranges and EWS parameters. Inpatient clinicians may find value in using these curves and percentile cut points to identify patients with vital signs that deviate from the ranges observed in this sample of hospitalized children. These new findings also invite new research to validate the use of vital sign percentiles in the development of evidence-based parameters for physiologic monitor alarms, inpatient EHR vital sign alerts, MET calling criteria, and EWSs.

ACKNOWLEDGMENTS

We thank William Nieczpiel, senior data architect at CHOP, and the research data warehouse team at CCHMC for their assistance obtaining the data.

REFERENCES


FIGURE 3

Scatterplot array showing the distribution of HR and RR in the study sample in comparison with existing EWS point ranges. Each point on the scatterplot represents 1 vital sign observation. For each score, observations that would score 0 points are colored black, and observations that would score 1, 2, 3, or 4 points are colored according to each scatter plot’s legend. We found that 14% to 38% of HR observations and 15% to 30% of respiratory rate observations would have resulted in increased total EWSs.


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